

Modeling tail dependence using Gaussian copula

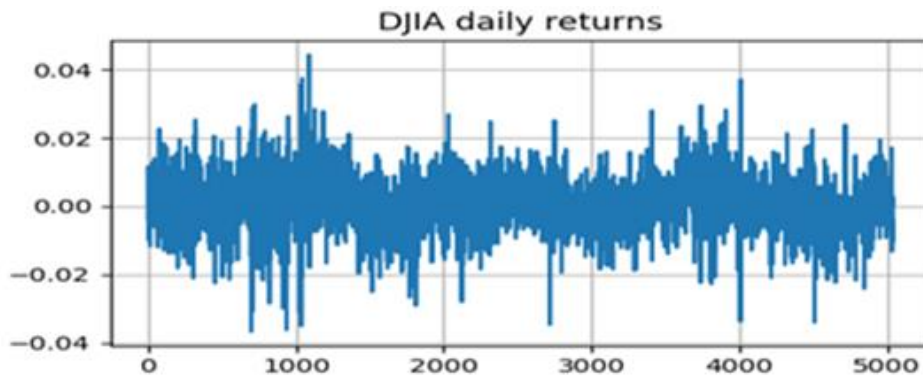
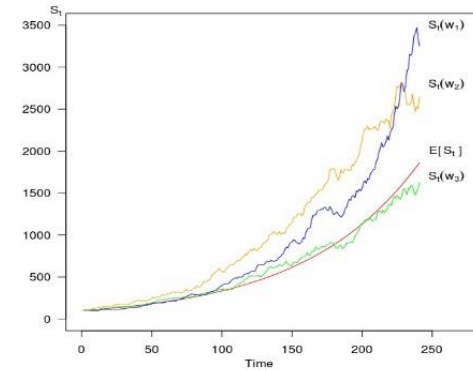
국립공주대학교 응용수학과
마 용 기
in collaboration with Ciprian Necula

2024년 11월 29일

- 확률미분방정식
- 의존관계 모형화
- 일반화된 섭동가우시안 코풀라
- 실증분석

확률미분방정식(I)

$$\frac{dS_t}{S_t} = \mu dt + \sigma dB_t \quad (\mu: \text{기대수익률}, \sigma: \text{변동성})$$



$$\sigma \rightarrow \sigma(Y_t) \rightarrow \sigma(Y_t, Z_t)$$

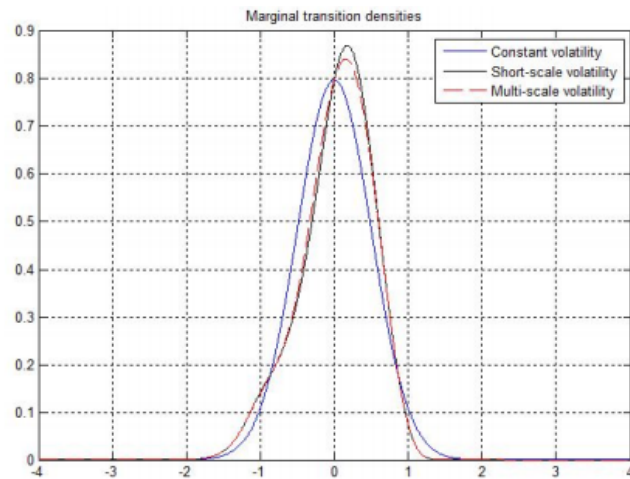
- $$\frac{dS_t}{S_t} = \mu dt + \sigma(Y_t, Z_t) dB_t$$

$$dY_t = \alpha_1(m_1 - Y_t)dt + \sigma_1 dB_t^{(Y)} \quad \sigma \rightarrow \sigma(Y_t) \rightarrow \sigma(Y_t, Z_t)$$

$$dZ_t = \alpha_2(m_2 - Z_t)dt + \sigma_2 dB_t^{(Z)}$$

- Fast scale volatility factor
: fast mean-reverting diffusion process
- Slow scale volatility factor
: slowly varying diffusion process

■ Marginal distribution

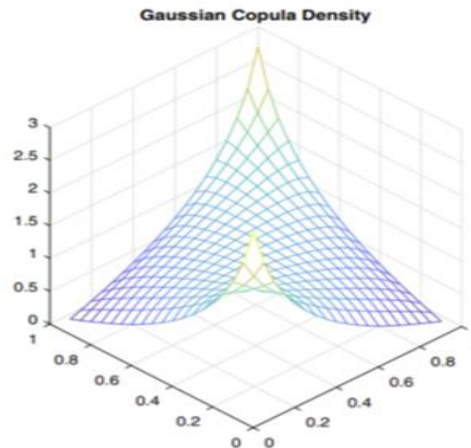


■ Linear correlation coefficient: : $-1 \leq \rho(X, Y) \leq 1$

- The variances of X and Y be must finite
- Independence of two random variables implies they are uncorrelated but zero correlation does not generally imply independence
- Linear correlation is not invariant under non-linear strictly increasing transformations

■ Copular function: Gaussian copula

$$\begin{aligned}dX_t^{(1)} &= \sigma_1 dB_t^{(1)} \\dX_t^{(2)} &= \sigma_2 dB_t^{(2)} \\dB_t^{(1)} dB_t^{(2)} &= \rho dt\end{aligned}$$



- Generally perturbed Gaussian copula

$$dX_t^{(1)} = f_1(Y_t, Z_t)dB_t^{(1)}$$

$$dX_t^{(2)} = f_2(Y_t, Z_t)dB_t^{(2)}$$

$$\text{with } dY_t = \alpha_1(m_1 - Y_t)dt + \sigma_1 dB_t^{(Y)}, \quad dZ_t = \alpha_2(m_2 - Z_t)dt + \sigma_2 dB_t^{(Z)}$$



$$u^{\epsilon, \delta} := \mathbb{P}\{X_T^{(1)} \in d\xi_1, X_T^{(2)} \in d\xi_2 | \mathbf{X}_t = \mathbf{x}, \mathbf{Y}_t = \mathbf{y}, \mathbf{Z}_t = \mathbf{z}\}$$

- Generally perturbed Gaussian copula

$$dX_t^{(1)} = f_1(Y_t, Z_t) dB_t^{(1)}$$

$$dX_t^{(2)} = f_2(Y_t, Z_t) dB_t^{(2)}$$

$$\text{with } dY_t = \alpha_1(m_1 - Y_t)dt + \sigma_1 dB_t^{(Y)}, \quad dZ_t = \alpha_2(m_2 - Z_t)dt + \sigma_2 dB_t^{(Z)}$$



$$dX_t^{(1)} = m e^{\alpha_1 Y_t + \beta_1 Z_t} dB_t^{(1)}$$

$$dX_t^{(2)} = m e^{\alpha_2 Y_t + \beta_2 Z_t} dB_t^{(2)}$$

$$\text{with } dY_t = \frac{1}{\epsilon} Y_t dt + \frac{\nu_1 \sqrt{2}}{\sqrt{\epsilon}} dB_t^{(Y)}, \quad dZ_t = -\delta Z_t dt + \frac{\nu_2 \sqrt{2}}{\sqrt{\epsilon}} dB_t^{(Z)}$$



$$u^{\epsilon, \delta}(t, x_1, x_2, z) \approx u_{0,0}(t, x_1, x_2, z) + \sqrt{\epsilon} u_{1,0}(t, x_1, x_2, z) + \sqrt{\delta} u_{0,1}(t, x_1, x_2, z)$$

- The leading-order term

$$u_{0,0} = \frac{1}{2\pi\bar{\sigma}_1(z)\bar{\sigma}_2(z)(T-t)\sqrt{1-\bar{\rho}^2(z)}} \\ \times \exp \left[-\frac{1}{2(1-\bar{\rho}^2(z))} \left\{ \frac{(\xi_1 - x_1)^2}{\bar{\sigma}_1^2(z)(T-t)} - 2\bar{\rho}(z) \frac{(\xi_1 - x_1)(\xi_2 - x_2)}{\bar{\sigma}_1(z)\bar{\sigma}_2(z)(T-t)} + \frac{(\xi_2 - x_2)^2}{\bar{\sigma}_2^2(z)(T-t)} \right\} \right]$$

- The first order correction terms

$$\sqrt{\epsilon}u_{1,0} = -(T-t) \left(R_1(z) \frac{\partial^3}{\partial x_1^3} + R_2(z) \frac{\partial^3}{\partial x_2^3} + R_{12}(z) \frac{\partial^3}{\partial x_1 \partial x_2^2} + R_{21}(z) \frac{\partial^3}{\partial x_1^2 \partial x_2} \right) u_{0,0}$$

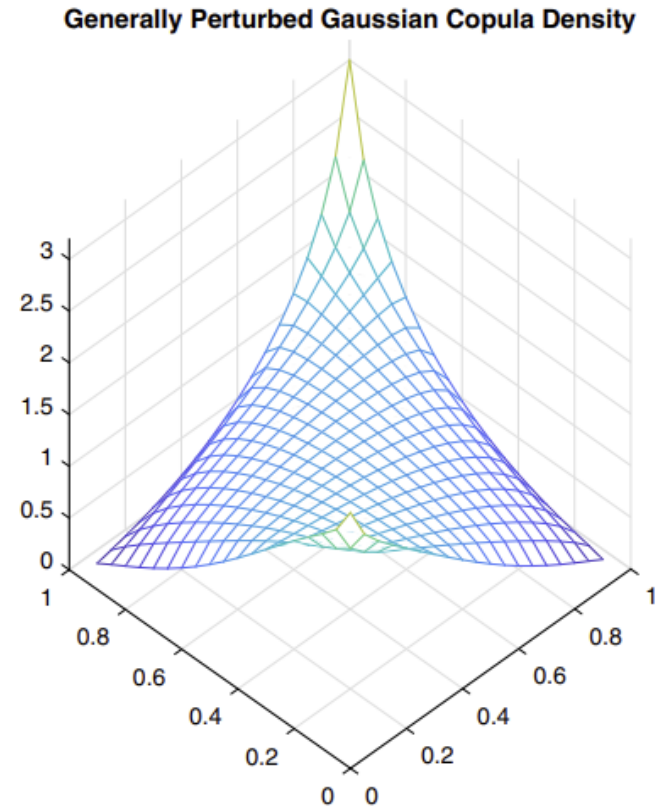
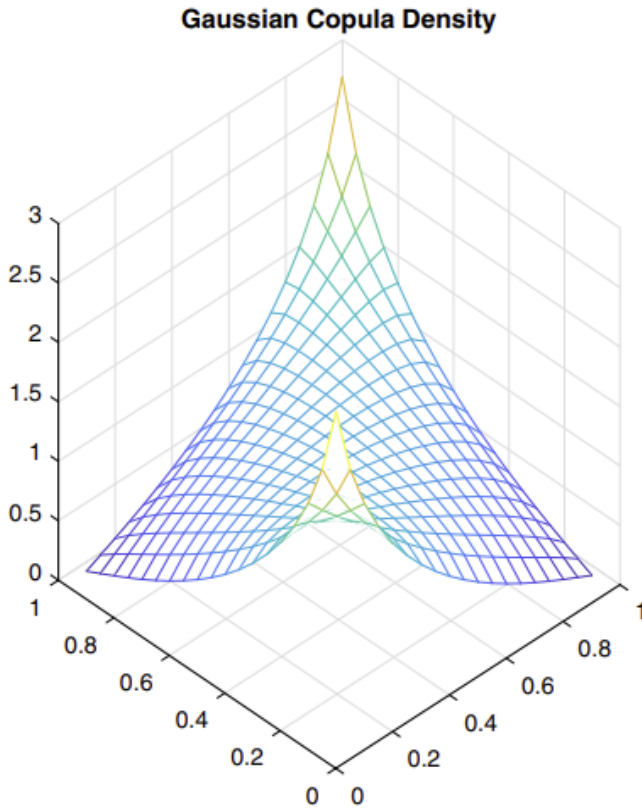
$$\sqrt{\delta}u_{0,1} = \frac{T-t}{2} \left(S_{12}(z) \frac{\partial^2}{\partial x_1 \partial z} + S_{21}(z) \frac{\partial^2}{\partial x_2 \partial z} \right) u_{0,0}$$

- Approximated transition density

$$\tilde{u}^{\epsilon,\delta} = \frac{1}{N} u_{0,0} \left(1 + \tanh \left(\sqrt{\epsilon} \frac{u_{1,0}}{u_{0,0}} + \sqrt{\delta} \frac{u_{0,1}}{u_{0,0}} \right) \right)$$

일반화된 섭동가우시안 코플라(IV)

- 가우시안 코플라와 일반화된 섭동가우시안 코플라의 밀도함수 비교



- 2000년부터 2017년까지 SPX지수와 DAX지수 데이터를 통한 모형의 모수 추정값

